

樣本空間

Sample Space and Events

- Sample Space Ω : the set of all possible outcomes in a random phenomenon. Examples:
 - 1. Sex of a newborn child: $\Omega = \{\text{girl, boy}\}$
 - 2. The order of finish in a race among the 7 horses 1, 2, ..., 7:

$$\Omega = \{\text{all } 7! \text{ Permutations of } (1, 2, 3, 4, 5, 6, 7)\}$$
 - 3. Flipping two coins: $\Omega = \{(H, H), (H, T), (T, H), (T, T)\}$
 - 4. Lifetime of a transistor: $\Omega = [0, \infty)$

known

outcome is not predictable

A ⊆ Ω

- Event: Any (measurable) subset of Ω is an event. Examples:

事件

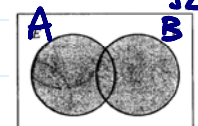
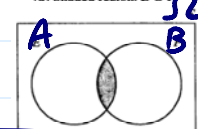
1. $A = \{\text{girl}\}$: the event - child is a girl.
 2. $A = \{\text{all outcomes in } \Omega \text{ starting with a 3}\}$: the event - horse 3 wins the race.

$$\Omega = \{\omega_1, \dots, \omega_n\}, \mathcal{Z} = \{\emptyset, \{\omega_1\}, \dots, \{\omega_n\}, \{\omega_1, \omega_2\}, \dots\}$$
 3. $A = \{(H, H), (H, T)\}$: the event - head appears on the 1st coin.
 4. $A = [0, 5]$: the event - transistor does not last longer than 5 hours.
- an event occurs: outcome \in the event 2^Ω : collection of all subsets in Ω .
 - **Q**: How many different events if $\#\Omega = n < \infty$? **Ans**: $2^n = \#2^\Omega$

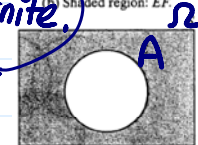
Set Operations of Events

p. 2-2

- Union. $C = A \cup B \Rightarrow C$: either A or B occurs
- Intersection. $C = A \cap B \Rightarrow C$: both A and B occur
- Complement. $C = A^c \Rightarrow C$: A does not occur
- Mutually Exclusive. $A \cap B = \emptyset \Rightarrow A$ and B have no outcomes in common. *disjoint.*

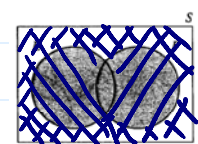
(a) Shaded region: $E \cup F$ (b) Shaded region: EF

- Definitions of union and intersection for more than two events can be defined in a similar manner

(c) Shaded region: E^c

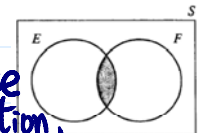
Some Simple Rules of Set Operations

- Commutative Laws. $A \cup B = B \cup A$ and $A \cap B = B \cap A$
- Associative Laws. $(A \cup B) \cup C = A \cup (B \cup C)$
 $(A \cap B) \cap C = A \cap (B \cap C)$
- Distributive Laws. $(A \cup B) \cap C = (A \cap C) \cup (B \cap C)$
 $(A \cap B) \cup C = (A \cup C) \cap (B \cup C)$
- DeMorgan's Laws.

(a) Shaded region: $E \cup F$

$$(\cup_{i=1}^n A_i)^c = \cap_{i=1}^n A_i^c \quad \text{and} \quad (\cap_{i=1}^n A_i)^c = \cup_{i=1}^n A_i^c$$

exercise + induction

(b) Shaded region: EF

Probability Measure 机率测度

• The Classical Approach

Sample Space Ω is a finite set

Probability: For an event A ,

$$P(A) = \frac{\#A}{\#\Omega}$$

a function: $2^\Omega \rightarrow [0,1]$

This explains why combinatorial thm play an important role in probability.

read more examples in textbook, sec. 2.5

Example (Roulette):

- $\Omega = \{0, 00, 1, 2, 3, 4, \dots, 35, 36\}$
- $P(\{\text{Red Outcome}\}) = 18/38 = 9/19$.

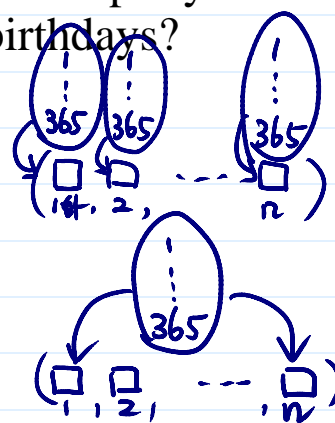


Example (Birthday Problem): n people gather at a party. What is the probability that they all have different birthdays?

- Ω = lists of n from $\{1, 2, 3, \dots, 365\}$
- A = {all permutations}
- $P_n(A) = (365)_n / 365^n$

n	8	16	24	32	40
$P_n(A)$.926	.716	.462	.247	.109

smaller



➤ Inadequacy of the Classical Approach

It requires:

- Finite Ω

- Symmetric Outcomes

Example (Sampling Proportional to Size):

- N invoices.
- Sample $n < N$.
- Pick large ones with higher probability.
- Note: Finite Ω , but non equally-likely outcomes.

$$P(A) = \frac{\#A}{\#\Omega}$$

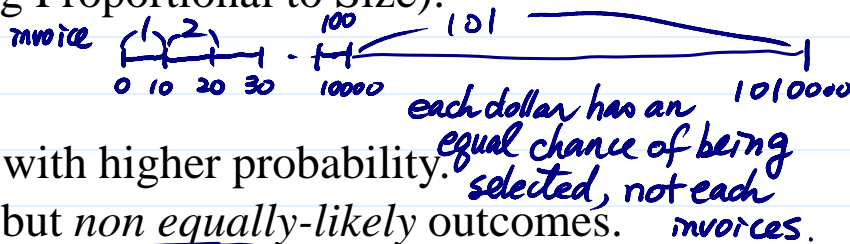
i.e. all outcomes in Ω are equally likely to occur

$$\omega \in \Omega, P(\omega) = \frac{1}{\#\Omega}$$

$N=101$
 $n=4$



$$\#\Omega = \binom{N}{n}$$



Example (Waiting for a success):

- Play roulette until a win.
- $\Omega = \{1, 2, 3, \dots\}$
- $P = ??$

discrete sample space
 $P: 2^\Omega \rightarrow [0,1]$

Example (Uniform Spinner):

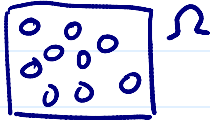
- Random Angle (in radians).
- $\Omega = (-\pi, \pi]$
- $P = ??$

continuous sample space
 $P: \sigma\text{-field} \rightarrow [0,1]$

• The Modern Approach

抽象化 A probability measure on Ω is a function P from subsets of Ω to the real number that satisfies the following **axioms** **公理**

- check if the 3 axioms hold on classical approach.**
- (Ax1) **Non-negativity**. For any event A , $P(A) \geq 0$. **- self-evident truth**
 - (Ax2) **Total one**. $P(\Omega) = 1$. **- a statement generally accepted as true.**
 - (Ax3) **Additivity**. If (A_1, A_2, \dots) is a sequence of **mutually exclusive events**, i.e., $A_i \cap A_j = \emptyset$ when $i \neq j$, then **countably many.**

$$P(A_1 \cup A_2 \cup \dots) = P(A_1) + P(A_2) + \dots$$


■ Notes:

- These axioms restrict probabilities, but do not **define** them.
- Probability is a property of events. **c.f.** $P(A) = \frac{\#A}{\#\Omega}$

➤ Define Probability Measures in a Discrete Sample Space.

- Q:** Is it required to define probabilities on every events? (e.g., n possible outcomes in Ω , $2^n - 1$ possible events) **$P(\Omega) = 1$.**

- Suppose $\Omega = \{\omega_1, \omega_2, \dots\}$, finite or countably infinite, let $p: \Omega \rightarrow R$ satisfy

$$p(\omega) \geq 0 \text{ for all } \omega \in \Omega \quad \text{and} \quad \sum_{\omega \in \Omega} p(\omega) = 1.$$

Let

$$P(A) = \sum_{\omega \in A} p(\omega)$$

for $A \subset \Omega$, then P is a probability measure. **(exercise)**

(Q: how to define p ?) **→ see the arguments about subjective & objective interpretation (LNp.2-15~17)**

- Example:** In the classical approach, $p(\omega) = 1/\#\Omega$. For example, throw a fair dice, $\Omega = \{1, \dots, 6\}$, $p(1) = \dots = p(6) = 1/6$ and $P(\text{odd}) = P(\{1, 3, 5\}) = p(1) + p(3) + p(5) = 3/6 = 1/2$.

- Example (non equally-likely events): Throwing an unfair dice might have $p(1) = 3/8$, $p(2) = p(3) = \dots = p(6) = 1/8$, and $P(\text{odd}) = P(\{1, 3, 5\}) = p(1) + p(3) + p(5) = 5/8$. (c.f., Example in LNp.2-4)

- Example (Waiting for Success – Play Roulette Until a Win):

- Q: what if we directly define P on 2^Ω ?** □ Let $r = 9/19$ and $q = 1 - r = 10/19$

- $\Omega = \{1, 2, 3, \dots\}$ **← infinite, countable.**

- Intuitively, $p(1) = r$, $p(2) = qr$, $p(3) = q^2r$, \dots , $p(n) = q^{n-1}r$, $\dots > 0$, and

$$\sum_{n=1}^{\infty} p(n) = \sum_{n=1}^{\infty} r q^{n-1} = \frac{r}{1-q} = 1.$$

► Proposition: If A_1, A_2, \dots, A_n are events in a sample space Ω , then

$$P(A_1 \cup \dots \cup A_n) \leq P(A_1) + \dots + P(A_n). \quad \text{cf } \textcircled{2} \text{ \& } \textcircled{4}$$

proof: $P((A_1 \cup \dots \cup A_{n-1}) \cup A_n) \leq P(A_1 \cup \dots \cup A_{n-1}) + P(A_n)$

by $\textcircled{4} \uparrow \leq P(A_1 \cup \dots \cup A_{n-1}) + P(A_{n-1}) + P(A_n) \leq \dots$

generalization of $\textcircled{4}$

► Proposition (inclusion-exclusion identity): If A_1, A_2, \dots, A_n are any n events, let

$$\sigma_1 = \sum_{i=1}^n P(A_i),$$

$$\sigma_2 = \sum_{1 \leq i < j \leq n} P(A_i \cap A_j),$$

$\rightarrow \binom{n}{2}$ different (i, j)

$$\sigma_3 = \sum_{1 \leq i < j < k \leq n} P(A_i \cap A_j \cap A_k),$$

$$\dots = \dots$$

$$\sigma_k = \sum_{1 \leq i_1 < \dots < i_k \leq n} P(A_{i_1} \cap \dots \cap A_{i_k})$$

$\rightarrow \binom{n}{k}$ different (i_1, \dots, i_k)

$$\dots = \dots$$

$$\sigma_n = P(A_1 \cap A_2 \cap \dots \cap A_n).$$

Q: For an outcome w contained in m out of the n events, how many times is its probability repetitively counted in $\sigma_1, \dots, \sigma_n$?

An intuitive proof for discrete sample space.

$P(w)$ counted m times in σ_1

$$\vdots \quad \binom{m}{2} \quad \vdots \quad \sigma_2$$

$$\vdots \quad \binom{m}{k} \quad \vdots \quad \sigma_k$$

$$\vdots \quad \vdots \quad \vdots$$

$$\vdots \quad \vdots \quad \vdots$$

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$$\vdots \quad \vdots \quad \vdots$$

$$P(A_1 \cup \dots \cup A_n) = \sigma_1 - \sigma_2 + \sigma_3 - \dots + (-1)^{k+1} \sigma_k + \dots + (-1)^{n+1} \sigma_n.$$

proof: prove by induction.

$n=2$, it holds \leftarrow by $\textcircled{4}$

$$n=3, P(\underline{A_1 \cup A_2 \cup A_3}) \stackrel{?}{=} P(A_1 \cup A_2) + P(A_3) - P(\overbrace{(A_1 \cup A_2) \cap A_3}^{(A_1 \cap A_3) \cup (A_2 \cap A_3)})$$

\ll by $\textcircled{4}$

\ll by $\textcircled{4}$

$$P(A_1) + P(A_2) - P(A_1 \cap A_2) + P(A_1 \cap A_3) + P(A_2 \cap A_3) - P(A_1 \cap A_2 \cap A_3)$$

$$= P(A_1) + P(A_2) + P(A_3) - P(A_1 \cap A_2) - P(A_1 \cap A_3) - P(A_2 \cap A_3) + P(A_1 \cap A_2 \cap A_3)$$

For $n > 3$, by mathematical induction.

■ Notes:

□ There are $\binom{n}{k}$ summands in σ_k

□ In symmetric examples, e.g. $k=2, P(A_1 \cap A_2) = P(A_{i_1} \cap A_{i_2}) \dots$

$$\sigma_k = \binom{n}{k} P(A_1 \cap \dots \cap A_k)$$

□ It can be shown that

for proof see textbook.

$$P(A_1 \cup \dots \cup A_n) \leq \sigma_1$$

$$P(A_1 \cup \dots \cup A_n) \geq \sigma_1 - \sigma_2$$

$$P(A_1 \cup \dots \cup A_n) \leq \sigma_1 - \sigma_2 + \sigma_3$$

$$\dots \dots \dots$$

■ Example (The Matching Problem).

□ Applications: (a) Taste Testing. (b) Gift Exchange.

□ Let Ω be all permutations $\omega = (i_1, \dots, i_n)$ of $1, 2, \dots, n$.

Thus, $\#\Omega = n!$.

□ Let

$$A_j = \{\omega: i_j = j\} \quad \text{and} \quad A = \bigcup_{i=1}^n A_i,$$

Q: $P(A) = ?$ (What would you expect when n is large?)

□ By symmetry,

$$\sigma_k = \binom{n}{k} P(A_1 \cap \dots \cap A_k),$$

□ Here,

$$P(A_1) = \frac{1 \times (n-1)!}{n!} = \frac{1}{n},$$

$$P(A_1 \cap A_2) = \frac{(n-2)!}{n!} = \frac{1}{(n)_2},$$

$$\dots = \dots,$$

$$P(A_1 \cap \dots \cap A_k) = \frac{(n-k)!}{n!} = \frac{1}{(n)_k}.$$

for $k = 1, \dots, n$.

□ So, $\sigma_k = \binom{n}{k} \frac{1}{(n)_k} = \frac{1}{k!},$

$$P(A) = \sigma_1 - \sigma_2 + \dots + (-1)^{n+1} \sigma_n = \sum_{k=1}^n (-1)^{k+1} \frac{1}{k!},$$

$$P(A) = 1 - \sum_{k=0}^n (-1)^k \frac{1}{k!} \approx 1 - \frac{1}{e} \approx 0.632 \Rightarrow P(A^c) \approx e^{-1} = 0.368$$

when $n \rightarrow \infty$

□ Note: approximation accurate to 3 decimal places if $n \geq 6$.

► Proportion: If A_1, A_2, \dots is a *partition* of Ω , i.e.,

$$1. \bigcup_{i=1}^{\infty} A_i = \Omega,$$

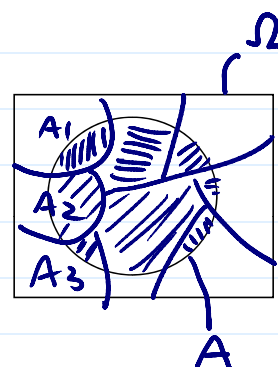
$$2. A_1, A_2, \dots, \text{ are mutually exclusive,}$$

then, for any event $A \subset \Omega$,

$$P(A) = \sum_{i=1}^{\infty} P(A \cap A_i).$$

proof: $A = A \cap \Omega = A \cap \left(\bigcup_{i=1}^{\infty} A_i \right) = \bigcup_{i=1}^{\infty} (A \cap A_i)$ mutually exclusive

$$P(A) = P\left(\bigcup_{i=1}^{\infty} (A \cap A_i) \right) = \sum_{i=1}^{\infty} P(A \cap A_i)$$



Monotone Sequences

p. 2-13

Q: How to define probability in a continuous sample space?

Definition: A sequence of events A_1, A_2, \dots , is called *increasing*

$$A_1 \subset A_2 \subset \dots \subset A_n \subset A_{n+1} \subset \dots \subset \Omega$$

and *decreasing* if

$$A_1 \supset A_2 \supset \dots \supset A_n \supset A_{n+1} \supset \dots \supset \emptyset$$

The limit of an increasing sequence is defined as

$$\lim_{n \rightarrow \infty} A_n = \bigcup_{i=1}^{\infty} A_i$$

and the limit of a decreasing sequence is

$$\lim_{n \rightarrow \infty} A_n = \bigcap_{i=1}^{\infty} A_i$$

Example: If $\Omega = \mathbb{R}$ and $A_k = (-\infty, 1/k)$, then A_k 's are decreasing and

$$\lim_{k \rightarrow \infty} A_k = \{\omega : \omega < 1/k \text{ for all } k \in \mathbb{Z}_+\} = (-\infty, 0].$$

Proportion: If A_1, A_2, \dots , is increasing or decreasing, then

De Morgan's Law: $\left(\lim_{n \rightarrow \infty} A_n\right)^c = \lim_{n \rightarrow \infty} A_n^c$

$$\left(\bigcup_{i=1}^{\infty} A_i\right)^c = \bigcap_{i=1}^{\infty} A_i^c = \lim_{i \rightarrow \infty} A_i^c$$

$$\left(\bigcap_{i=1}^{\infty} A_i\right)^c = \bigcup_{i=1}^{\infty} A_i^c = \lim_{i \rightarrow \infty} A_i^c$$

p. 2-14

Proportion: If A_1, A_2, \dots , is increasing or decreasing, then

satisfy (Ax1)-(Ax3) $\lim_{n \rightarrow \infty} P(A_n) = P\left(\lim_{n \rightarrow \infty} A_n\right)$ (5) $B_i \cap B_j = \emptyset$

proof (1) A_n increasing

Let $B_n = A_n \cap A_{n-1}^c \Rightarrow \bigcup_{n=1}^{\infty} B_n = \bigcup_{n=1}^{\infty} A_n$, $\bigcap_{n=1}^{\infty} B_n = \bigcap_{n=1}^{\infty} A_n$

$$P\left(\lim_{n \rightarrow \infty} A_n\right) = P\left(\bigcup_{n=1}^{\infty} A_n\right) = P\left(\bigcup_{n=1}^{\infty} B_n\right) = \sum_{n=1}^{\infty} P(B_n) = \lim_{k \rightarrow \infty} \sum_{n=1}^k P(B_n)$$

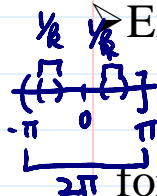
(2) A_n decreasing, $\Rightarrow A_n^c$ increasing.

$$1 - P\left(\lim_{n \rightarrow \infty} A_n\right) = 1 - P\left(\bigcap_{n=1}^{\infty} A_n\right) = P\left(\left(\bigcap_{n=1}^{\infty} A_n\right)^c\right)$$

$$= P\left(\bigcup_{n=1}^{\infty} A_n^c\right) = P\left(\lim_{n \rightarrow \infty} A_n^c\right) = \lim_{n \rightarrow \infty} P(A_n^c)$$

$$= \lim_{n \rightarrow \infty} 1 - P(A_n) = 1 - \lim_{n \rightarrow \infty} P(A_n)$$

Example (Uniform Spinner): Let $\Omega = (-\pi, \pi]$. Define



c.f. p.m. for discrete sample space

$$P((a, b]) = \frac{b-a}{2\pi}$$

continuous sample space

play a role similar to

w_1, w_2, \dots in small p for discrete

for subintervals $(a, b] \subset \Omega$. Then, extend P to other subsets using the 3 axioms. For example, if $-\pi < a < b < \pi$,

sample space

$$P([a, b]) = P\left(\left(\bigcap_{k=1}^{\infty} \left(a - \frac{1}{k}, b\right] \cap \Omega\right)\right) = P\left(\bigcap_{k=1}^{\infty} \left(\left(a - \frac{1}{k}, b\right] \cap \Omega\right)\right)$$

Note $[a, b] \neq (a, b]$
 $\Rightarrow P$ not defined
 in previous slide.

$$= \lim_{k \rightarrow \infty} P\left(\left(a - \frac{1}{k}, b\right] \cap \Omega\right)$$

$$= \lim_{k \rightarrow \infty} \frac{1}{2\pi} \left(b - a + \frac{1}{k}\right) = \frac{b - a}{2\pi}$$

$\left[\left(a - \frac{1}{k}, b\right] \cap \Omega\right)$

$$A_k = \left(a - \frac{1}{k}, b\right] \downarrow \lim_{k \rightarrow \infty} A_k = [a, b]$$

Some notes

$$\square P(\{a\}) = P([a, b] - (a, b]) = P([a, b]) - P((a, b]) = 0.$$

\square If $C = \{\omega_1, \omega_2, \dots\} \subset \Omega$, then

$$P(C) = \sum_{i=1}^{\infty} P(\{\omega_i\}) = 0 + 0 + \dots = 0.$$

$\{\omega_i\}$

a countable set.

\square The probability of a rational outcome is zero

Objective vs. Subjective Interpretations of Probability

➤ **Q:** What do we mean if we say that the probability of rain tomorrow is 40%? \leftarrow how to interpret it?

Objective: Long run relative frequencies

Subjective: Chosen to reflect opinion

➤ The Objective (Frequency) Interpretation 頻率

■ Through Experiment: Imagine the experiment repeated N times. For an event A , let

$N_A = \#$ occurrences of A .

Then,

$$P(A) \equiv \lim_{N \rightarrow \infty} \frac{N_A}{N}$$

frequency

■ Example (Coin Tossing):

N	100	1000	10000	100000
N_H	55	493	5143	50329
N_H/N	.550	.493	.514	.503

$\rightarrow 0.5$

The result is consistent with $P(H)=0.5$.

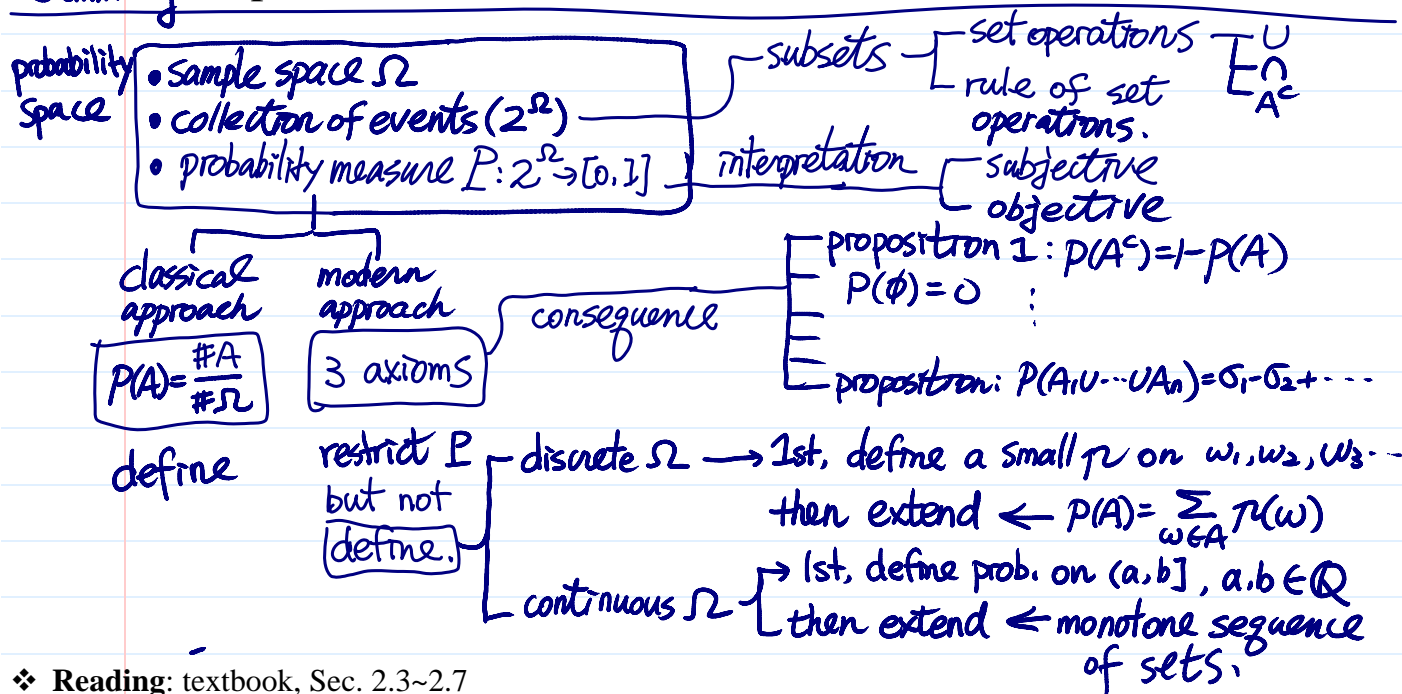
odds of an event \rightarrow 勝
 $O(A) = \frac{P(A)}{P(A^c)} = \frac{P(A)}{1-P(A)}$ 敗
 率

➤ The Subjective Interpretation

Strategy: Assess probabilities by imagining bets
 Examples:
 $O(A) = \frac{P(A)}{1-P(A)} \geq \frac{1}{2}$
 $P(A) \geq \frac{O(A)}{1+O(A)}$
 Peter is willing to give two to one odds that it will rain tomorrow. His subjective probability for rain tomorrow is at least $\frac{2}{3}$

$O(A^c) = \frac{P(A^c)}{P(A)} = \frac{1-P(A)}{P(A)}$ Paul accepts the bet. His subjective probability for rain tomorrow is at most $\frac{2}{3}$ present personal degree of belief.
 $\Rightarrow P(A) \leq \frac{2}{3}$ Probabilities are simply personal measures of how likely we think it is that a certain event will occur
 e.g. how likely you think Shakespeare wrote Hamlet?
 This can be applied even when the idea of repeated experiments is not feasible

Summary



❖ Reading: textbook, Sec. 2.3~2.7